

Questions 2 and 7 will be marked. In all of the below, assume that any design matrices  $X$  are  $n \times p$  and have their columns centred and then scaled to have  $\ell_2$ -norm  $\sqrt{n}$ .

1. In the setting of Theorem 23, assume instead of the compatibility condition that for some  $c \in (0, 1)$  there exists  $\phi > 0$  such that for all  $\delta \in \mathbb{R}^p$  with  $(1 - c)\|\delta_N\|_1 \leq (1 + c)\|\delta_S\|_1$ , we have

$$\|\delta_S\|_1^2 \leq \frac{s\|X\delta\|_2^2}{n\phi^2}.$$

Let  $\hat{\beta}$  be a Lasso estimator with regularisation parameter  $\lambda > 0$ . Show that with probability at least  $1 - 2p^{-(c^2 A^2/2-1)}$ , we have both

$$\frac{1}{n}\|X(\hat{\beta} - \beta^0)\|_2^2 \leq (1 + c)^2 \lambda^2 \frac{s}{\phi^2} \quad \text{and} \quad \|\hat{\beta} - \beta^0\|_1 \leq \lambda \frac{2(1 + c)}{1 - c} \frac{s}{\phi^2}.$$

[Hint: Start with the improved version of the basic inequality from Qu. 11 of Sheet 2.]

2. Let  $Y = \mu^0 \mathbf{1} + X\beta^0 + \varepsilon$  and let  $S = \{k : \beta^0 \neq 0\}$ ,  $N := \{1, \dots, p\} \setminus S$ . Without loss of generality assume  $S = \{1, \dots, |S|\}$ . Assume that  $X_S$  has full column rank and let  $\Omega = \{\|X^\top \varepsilon\|_\infty/n \leq \lambda_0\}$ . Show that, when  $\lambda > \lambda_0$ , if the following two conditions hold

$$\begin{aligned} \sup_{\tau: \|\tau\|_\infty \leq 1} \|X_N^\top X_S (X_S^\top X_S)^{-1} \tau\|_\infty &< \frac{\lambda - \lambda_0}{\lambda + \lambda_0} \\ (\lambda + \lambda_0) \|\{(\frac{1}{n} X_S^\top X_S)^{-1}\}_k\|_1 &< |\beta_k^0| \quad \text{for } k \in S, \end{aligned}$$

then on  $\Omega$ , there exists a Lasso solution that satisfies  $\text{sgn}(\hat{\beta}_\lambda^L) = \text{sgn}(\beta^0)$ . Show moreover that the Lasso solution is unique.

3. Consider the setup of Question 1 with  $c = 1/2$  and write  $\hat{S} := \{j : \hat{\beta}_j \neq 0\}$  and set  $\hat{s} := |\hat{S}|$ .

(a) Show that on the event  $\Omega$ , for any non-empty subset  $B$  of  $\hat{S}$ , we have

$$\frac{1}{n} \text{sgn}(\hat{\beta}_B)^\top X_B^\top X(\beta^0 - \hat{\beta}) \geq \frac{\lambda|B|}{2}.$$

[Hint: Start with the KKT conditions.]

(b) Let  $\kappa_m^2$  be the maximum eigenvalue of  $X_M^\top X_M/n$  over all  $M \subset \{1, \dots, p\}$  with  $|M| \leq m$ . Prove that on  $\Omega$ , any  $B \subseteq \hat{S}$  satisfies

$$|B| \leq 9s\kappa_{|B|}^2/\phi^2.$$

Let

$$m^* = \min\{m \geq 1 : m > 9\kappa_m^2 s/\phi^2\},$$

with  $m^* = \infty$  if there does not exist any  $m$  satisfying the condition defining the set above. Deduce that on  $\Omega$ ,  $\hat{s} < m^*$ , and moreover that  $\hat{s} \leq 9\kappa_{m^*}^2 s/\phi^2$ .

4. (a) Show that

$$\max_{\theta: \|X^\top \theta\|_\infty \leq \lambda} G(\theta) = \frac{1}{2n} \|Y - X\hat{\beta}_\lambda^L\|_2^2 + \lambda \|\hat{\beta}_\lambda^L\|_1,$$

where

$$G(\theta) = \frac{1}{2n} \|Y\|_2^2 - \frac{1}{2n} \|Y - n\theta\|_2^2.$$

Show that the unique  $\theta$  maximising  $G$  is  $\theta^* = (Y - X\hat{\beta}_\lambda^L)/n$ . [Hint: Treat the Lasso optimisation problem as minimising  $\|Y - z\|_2^2/(2n) + \lambda\|\beta\|_1$  subject to  $z - X\beta = 0$  over  $(\beta, z) \in \mathbb{R}^p \times \mathbb{R}^n$  and consider the Lagrangian.]

(b) Let  $\tilde{\theta}$  be such that  $\|X^\top \tilde{\theta}\|_\infty \leq \lambda$ . Explain why if

$$\max_{\theta: G(\theta) \geq G(\tilde{\theta})} |X_k^\top \theta| < \lambda,$$

then we know that  $\hat{\beta}_{\lambda, k}^L = 0$ . By considering  $\tilde{\theta} = Y\lambda/(n\lambda_{\max})$  with  $\lambda_{\max} = \|X^\top Y\|_\infty/n$ , show that  $\hat{\beta}_{\lambda, k}^L = 0$  if

$$\frac{1}{n} |X_k^\top Y| < \lambda - \frac{\|Y\|_2}{\sqrt{n}} \frac{\lambda_{\max} - \lambda}{\lambda_{\max}}.$$

5. Suppose  $\hat{\beta}$  is a square-root Lasso solution from a regression of  $Y$  onto  $X$  with regularisation parameter  $\gamma > 0$ . Show that provided  $Y - \bar{Y}\mathbf{1} \neq X\hat{\beta}$ , we have

$$\frac{1}{\sqrt{n}} \frac{(Y - \bar{Y}\mathbf{1})^\top (Y - \bar{Y}\mathbf{1} - X\hat{\beta})}{\|Y - \bar{Y}\mathbf{1} - X\hat{\beta}\|_2} = \frac{1}{\sqrt{n}} \|Y - \bar{Y}\mathbf{1} - X\hat{\beta}\|_2 + \gamma \|\hat{\beta}\|_1.$$

6. The elastic net estimator in the linear model minimises

$$\frac{1}{2n} \|Y - X\beta\|_2^2 + \lambda(\alpha\|\beta\|_1 + (1 - \alpha)\|\beta\|_2^2/2)$$

over  $\beta \in \mathbb{R}^p$ , where  $\alpha \in [0, 1]$  is fixed.

(a) Suppose  $X$  has two columns  $X_j$  and  $X_k$  that are identical and  $\alpha < 1$ . Explain why the minimising  $\beta^*$  above is unique and has  $\beta_k^* = \beta_j^*$ .  
 (b) Let  $\hat{\beta}^{(0)}, \hat{\beta}^{(1)}, \dots$  be the solutions from iterations of a coordinate descent procedure to minimise the elastic net objective. For a fixed variable index  $k$ , let  $A = \{1, \dots, k-1\}$  and  $B = \{k+1, \dots, p\}$  and find the form of  $\hat{\beta}_k^{(m)}$  in terms of  $\hat{\beta}_A^{(m)}$  and  $\hat{\beta}_B^{(m-1)}$ .

7. Consider the model

$$Y = X(\beta^0 + \delta^0) + \varepsilon,$$

where  $\delta^0 \in \mathbb{R}^p$  represents a dense perturbation of the usual sparse linear model defined by  $\beta^0 \in \mathbb{R}^p$  alone. The *Lava estimator*  $(\hat{\beta}_\lambda, \hat{\delta}_\lambda)$  with tuning parameter  $\lambda = (\lambda_1, \lambda_2)^\top \in (0, \infty)^2$  is defined by

$$(\hat{\beta}_\lambda, \hat{\delta}_\lambda) := \underset{(\beta, \delta) \in \mathbb{R}^p \times \mathbb{R}^p}{\operatorname{argmin}} \left\{ \frac{1}{2n} \|Y - X(\beta + \delta)\|_2^2 + \lambda_1 \|\beta\|_1 + \lambda_2 \|\delta\|_2^2 \right\}.$$

Find an expression for  $\hat{\delta}_\lambda$  involving  $X$ ,  $Y$ ,  $\hat{\beta}_\lambda$ ,  $\lambda_2$  and  $n$ . Deduce that  $\hat{\beta}_\lambda$  is the minimiser of a Lasso objective with transformed design matrix  $\tilde{X} := AX$  and transformed response  $AY$ , where  $A := (I - XQX^\top)^{1/2}$  and  $Q := (X^\top X + 2n\lambda_2 I)^{-1} \in \mathbb{R}^{p \times p}$ .

Now let  $\Omega := \{\|\tilde{X}^\top(\tilde{X}\delta^0 + A\varepsilon)\|_\infty/n \leq \lambda_1\}$ . Show that on  $\Omega$ , we have

$$\frac{1}{n} \|\tilde{X}(\hat{\beta}_\lambda - \beta^0)\|_2^2 \leq 4\lambda_1 \|\beta^0\|_1.$$

Conclude that on  $\Omega$  we have

$$\frac{1}{n} \|X(\hat{\beta}_\lambda - \beta^0)\|_2^2 \leq \frac{4\lambda_1 \|\beta^0\|_1}{1 - \kappa}$$

when the maximum eigenvalue of  $XQX^\top$  is  $\kappa < 1$ .

8. By using applications of the weak union and contraction properties or otherwise, answer the following.

(a) Suppose  $X_1, X_2, \dots$  and  $Y_1, Y_2, \dots$  are sequences of random vectors satisfying

$$Y_t \perp\!\!\!\perp X_{t-1}, \dots, X_1 \mid Y_{t-1}, \dots, Y_1$$

for all  $t \in \mathbb{N}$ . Show that for all  $r \in \mathbb{N}$ , we also have

$$Y_t, \dots, Y_{t+r} \perp\!\!\!\perp X_{t-1}, \dots, X_1 \mid Y_{t-1}, \dots, Y_1.$$

(b) Suppose  $(X_i, Y_i, Z_i)_{i=1}^n$  are independent triples satisfying  $X_i \perp\!\!\!\perp Y_i \mid Z_i$ . Show that

$$X_1, \dots, X_n \perp\!\!\!\perp Y_1, \dots, Y_n \mid Z_1, \dots, Z_n.$$

[Hint: Argue that it suffices to show the result for  $n = 2$ .]

9. Let  $Z \sim N_p(\mu, \Sigma)$  with  $\Sigma$  positive definite. Show that for any  $A, B \subset [p]$ ,

$$Z_A \mid Z_B = z_B \sim N_{|A|}(\mu_A + \Sigma_{A,B} \Sigma_{B,B}^{-1} (z_B - \mu_B), \Sigma_{A,A} - \Sigma_{A,B} \Sigma_{B,B}^{-1} \Sigma_{B,A}).$$

Here  $\Sigma_{A,B}$ , for example, is the submatrix of  $\Sigma$  formed of those rows and columns indexed by  $A$  and  $B$  respectively. [Hint: Find a matrix  $M \in \mathbb{R}^{|A| \times |B|}$  such that  $Z_A - MZ_B$  and  $Z_B$  are independent.]