

Gaussian mixtures: convexity properties and CLT rates for the entropy and Fisher information

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Outline

- 1 Entropy
- 2 Fisher Information
- 3 CLT rates
- 4 An Open Question

X_1, X_2 independent RVs

- Entropy Power Inequality (EPI)

$$h(\sqrt{1-t}X_1 + \sqrt{t}X_2) \geq (1-t)h(X_1) + th(X_2)$$

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Q: Where is it maximized?

Entropy

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For X_1, X_2 i.i.d. log-concave

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- ▶ EPI special case
- ▶ **No** special case **known** (of X_1, X_2)!

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- $g_X \propto e^{-|x|^p}$ (p- stable, $p \in (0, 2]$)
- Cauchy $\propto \frac{1}{\pi(1+x^2)} \Rightarrow \frac{Z_1}{|Z_2|}$

Entropy: Gaussian Mixtures

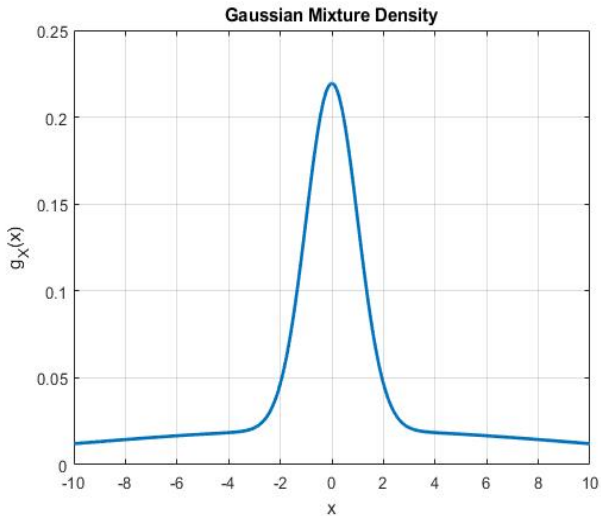


Figure: $\mathbb{P}(Y = 1) = \mathbb{P}(Y = 10) = 0.5$

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- ▶ In [Costa '85]: strengthened EPI for **stable laws**

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Cramér-Rao (C-R):

$$I(X) \geq \frac{1}{\text{Var}(X)} \quad (d = 1)$$

Fisher Information

Proposition

Let $X = Y \cdot Z$ a *GM*, $\text{Var}(X) = \mathbb{E}[Y^2]$

$$\frac{1}{\mathbb{E}[Y^2]} \stackrel{\text{C-R}}{\leq} I(X) \leq \mathbb{E}\left[\frac{1}{Y^2}\right]$$

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$$I(X) = \int \frac{(f'(x))^2}{f(x)} dx \leq \int x^2 \mathbb{E}\left[\frac{1}{Y^4\sqrt{2\pi}Y^2}e^{-\frac{x^2}{2Y^2}}\right] dx = \mathbb{E}\left[\frac{1}{Y^2}\right] \quad \square$$

III. Fisher Information

(GM) in $\mathbb{R}^d$: $X = \mathbf{Y} \cdot Z$,
 $Z \sim \mathcal{N}(0, \mathbf{I}_{d \times d})$, $\mathbf{Y} \in \mathbb{R}^{d \times d}$, symmetric, $\mathbb{P}(\mathbf{Y} \succeq 0) = 1$

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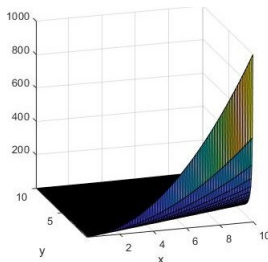
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Proposition (Bobkov, '22)

The Fisher information satisfies “Jensen’s” inequality: let π be a probability measure on the space of probability densities in $\mathbb{R}^n$. Then, if f is any mixture of densities $f = \int g\pi(dg)$,

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Proposition

The Fisher information matrix is convex as a matrix-valued functional and satisfies “Jensen’s” inequality: if f is any mixture of densities $f = \int g\pi(dg)$,

$$\mathcal{I}(f) \preceq \int \mathcal{I}(g)\pi(dg)$$

provided that $\int \|\mathcal{I}(g)\|_{\text{OP}} < \infty$

Proof. $R(x, \lambda) = \frac{xx^T}{\lambda}$ is jointly operator convex + Bobkov's proof with $\langle \mathcal{I}(X)x, x \rangle$ □

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$\Rightarrow$

$$\mathcal{I}(\mathbf{Y}Z) \preceq \mathbb{E}_{\mathbf{Y}} \mathcal{I}(Z_{\mathbf{Y}}) = \mathbb{E}(\mathbf{Y}\mathbf{Y}^T)^{-1},$$

where $Z_{\Sigma} \sim \mathcal{N}(0, \Sigma\Sigma^T)$

$(a_1, \dots, a_n) : \sum_{i=1}^n a_i^2 = 1, X_i = Y_i Z_i$ GMs

$$S_n = \sum_{i=1}^n a_i X_i = \sum_{i=1}^n a_i Y_i Z_i \stackrel{d}{=} \left(\sum_{i=1}^n a_i^2 Y_i^2 \right)^{1/2} Z$$

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Standardised Fisher information:

$$J_{st}(S_n) := \text{Var}(S_n) I(S_n) - 1 \leq (\mathbb{E} Y_1^2) \mathbb{E} \left[\frac{1}{\sum_{i=1}^n a_i^2 Y_i^2} \right] - 1$$

Corollary

Suppose $\mathbb{E}Y_1^{2+2\delta}, \mathbb{E}Y_1^{-2-2\delta} < \infty$. Then

$$J_{st}(S_n) \leq C_{Y_1} \|a\|_{\frac{2\delta}{2+2\delta}}^{\frac{2\delta}{(1+\delta)}}$$

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- E.g. $a_i = \frac{1}{\sqrt{n}}$, then

$$\|a\|_{2+2\delta}^{\frac{2\delta}{1+\delta}} = \left(\sum_{i=1}^n \frac{1}{n^{1+\delta}} \right)^{\frac{2\delta}{(1+\delta)(2+2\delta)}} = \left(\frac{1}{n} \right)^{\frac{\delta^2}{(1+\delta)^2}} \rightarrow 0$$

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- $O(\|a\|_4^4)$ in [Artstein, Ball, Barthe, Naor, '04] for weighted sums under **much stronger moment assumptions**

Corollary (General dimension)

If $\mathbb{E}\|\mathbf{Y}\mathbf{Y}^T\|_{\text{OP}}^{1+\delta} < \infty$ and $\mathbb{E}\|(\mathbf{Y}\mathbf{Y}^T)^{-1}\|_{\text{OP}}^{1+\delta} < \infty$,

$$\|\text{Cov}(S_n)^{\frac{1}{2}}\mathcal{I}(S_n)\text{Cov}(S_n)^{\frac{1}{2}} - \text{I}_d\|_{\text{OP}} \leq C(\mathbf{Y}) \log^\delta(d+1) \|a\|_{\frac{2+2\delta}{1+\delta}}^{\frac{2\delta}{1+\delta}}$$

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- ▶ Convergence of the Fisher information **matrix**
- ▶ **First** such result!

An Open Question

► Is

$$I\left(\sqrt{t}X_1 + \sqrt{1-t}X_2\right)$$

minimized at $t = \frac{1}{2}$? More generally, is

$$I\left(\sum_{i=1}^n a_i X_i\right)$$

minimized at $a_i = \frac{1}{\sqrt{n}}$?

Thank you!

