## Mathematical Tripos: Part IB

## Statistics: Example Sheet 3 (of 3)

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- 1. (a) Let  $\mathbf{X} \sim N_n(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ , and let A be an arbitrary  $m \times n$  matrix. Prove directly from the definition that  $A\mathbf{X}$  has an m-variate normal distribution. Show that  $\operatorname{cov}(A\mathbf{X}) = A\boldsymbol{\Sigma}A^T$ , and that  $A\mathbf{X} \sim N_m(A\boldsymbol{\mu}, A\boldsymbol{\Sigma}A^T)$ . Give an alternative proof that  $A\mathbf{X} \sim N_m(A\boldsymbol{\mu}, A\boldsymbol{\Sigma}A^T)$  using moment generating functions.
  - (b) Let  $\mathbf{X} \sim N_n(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ , and let  $\mathbf{X}_1$  denote the first  $n_1$  components of  $\mathbf{X}$ . Let  $\boldsymbol{\mu}_1$  denote the first  $n_1$  components of  $\boldsymbol{\mu}$ , and let  $\boldsymbol{\Sigma}_{11}$  denote the upper left  $n_1 \times n_1$  block of  $\boldsymbol{\Sigma}$ . Show that  $\mathbf{X}_1 \sim N_{n_1}(\boldsymbol{\mu}_1, \boldsymbol{\Sigma}_{11})$ .
- 2. Let  $X_1, \ldots, X_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$ , where  $\sigma^2$  is unknown, and suppose we are interested in testing  $H_0: \mu = \mu_0$  against  $H_1: \mu \neq \mu_0$ . Letting  $\bar{X} = n^{-1} \sum_{i=1}^n X_i$  and  $S_{XX} = \sum_{i=1}^n (X_i \bar{X})^2$ , show that the likelihood ratio can be expressed as

$$\Lambda_{\mathbf{X}}(H_0, H_1) = \left(1 + \frac{T^2}{n-1}\right)^{n/2},$$

where  $T = \frac{n^{1/2}(\bar{X}-\mu_0)}{\{S_{XX}/(n-1)\}^{1/2}}$ . Determine the distribution of T under  $H_0$ , and hence determine the size  $\alpha$  likelihood ratio test.

- 3. Statisticians A and B obtain independent samples  $X_1, \ldots, X_{10}$  and  $Y_1, \ldots, Y_{17}$  respectively, both from a  $N(\mu, \sigma^2)$  distribution with both  $\mu$  and  $\sigma^2$  unknown. They estimate  $(\mu, \sigma^2)$  by  $(\bar{X}, S_{XX}/9)$  and  $(\bar{Y}, S_{YY}/16)$  respectively, where, for example,  $\bar{X} = \frac{1}{10} \sum_{i=1}^{10} X_i$  and  $S_{XX} = \sum_{i=1}^{10} (X_i \bar{X})^2$ . Given that  $\bar{X} = 5.5$  and  $\bar{Y} = 5.8$ , which statistician's estimate of  $\sigma^2$  is more probable to have exceeded the true value by more than 50%? Find this probability (approximately) in each case. [Hint: This is something of a 'trick' question. Why? You may find  $\chi^2$  tables helpful.]
- 4. Suppose that  $X_1, \ldots, X_m$  are iid  $N(\mu_X, \sigma_X^2)$ , and, independently,  $Y_1, \ldots, Y_n$  are iid  $N(\mu_Y, \sigma_Y^2)$ , with  $\mu_X, \mu_Y, \sigma_X^2$  and  $\sigma_Y^2$  unknown. Write down the distributions of  $S_{XX}/\sigma_X^2$  and  $S_{YY}/\sigma_Y^2$ . Derive a 100(1  $\alpha$ )% confidence interval for  $\sigma_X^2/\sigma_Y^2$ .
- 5. Consider the simple linear regression model  $Y_i = a + bx_i + \varepsilon_i$ , i = 1, ..., n, where  $\varepsilon_1, \ldots, \varepsilon_n \stackrel{iid}{\sim} N(0, \sigma^2)$  and  $\sum_{i=1}^n x_i = 0$ . Derive from first principles explicit expressions for the MLEs  $\hat{a}$ ,  $\hat{b}$  and  $\hat{\sigma}^2$ . Show that we can obtain the same expressions if we regard the simple linear regression model as a special case of the general linear model  $\mathbf{Y} = X\boldsymbol{\beta} + \boldsymbol{\epsilon}$  and specialise the formulae  $\hat{\boldsymbol{\beta}} = (X^T X)^{-1} X^T \mathbf{Y}$  and  $\hat{\sigma}^2 = n^{-1} ||\mathbf{Y} X\hat{\boldsymbol{\beta}}||^2$ .

6. Consider the model  $Y_i = bx_i + \varepsilon_i$ , i = 1, ..., n, where the  $\varepsilon_i$  are independent with mean zero and variance  $\sigma^2$  (regression through the origin). Write this in the form  $\mathbf{Y} = X\boldsymbol{\beta} + \boldsymbol{\varepsilon}$ , and find the least squares estimator of b.

The relationship between the range in metres, Y, of a howitzer with muzzle velocity v metres per second fired at angle of elevation  $\alpha$  degrees is assumed to be  $Y = \frac{v^2}{g}\sin(2\alpha) + \varepsilon$ , where g = 9.81 and where  $\varepsilon$  has mean zero and variance  $\sigma^2$ . Estimate v from the following independent observations made on 9 shells.

$\alpha ~(\text{deg})$	5	10	15	20	25	30	35	40	45
$\sin 2lpha$	0.1736	0.3420	0.5	0.6428	0.7660	0.8660	0.9397	0.9848	1
range (m)	4860	9580	14080	18100	21550	24350	26400	27700	28300

- 7. Consider the model  $Y_i = \mu + \varepsilon_i$ , i = 1, ..., n, where  $\varepsilon_i$  are iid  $N(0, \sigma^2)$  random variables. Write this in matrix form  $\mathbf{Y} = X\boldsymbol{\beta} + \boldsymbol{\varepsilon}$ , and find the MLE  $\hat{\boldsymbol{\beta}}$ . Find the fitted values, the residuals and the residual sum of squares. Show how applying Theorem 13.2 (in lectures) to this case gives the independence of  $\bar{Y}$  and  $S_{YY}$  for an iid sample from  $N(\mu, \sigma^2)$ . Write down an unbiased estimate  $\tilde{\sigma}^2$  of  $\sigma^2$ .
- 8. Consider the one-way analysis of variance (ANOVA) model  $Y_{ij} = \mu_i + \varepsilon_{ij}$ ,  $i = 1, \ldots, I$ ,  $j = 1, \ldots, n_i$ , where  $(\varepsilon_{ij}) \stackrel{iid}{\sim} N(0, \sigma^2)$ . Derive from first principles explicit expressions for the MLEs  $\hat{\mu}_1, \ldots, \hat{\mu}_I$  and  $\hat{\sigma}^2$ . Show that we can obtain the same expressions if we regard the ANOVA model as a special case of the general linear model  $\mathbf{Y} = X\boldsymbol{\beta} + \boldsymbol{\varepsilon}$ and specialise the formulae  $\hat{\boldsymbol{\beta}} = (X^T X)^{-1} X^T \mathbf{Y}$  and  $\hat{\sigma}^2 = n^{-1} \|\mathbf{Y} - X\hat{\boldsymbol{\beta}}\|^2$ .
- 9. Consider the linear model **Y** = Xβ + ε where **Y** is an n × 1 vector of observations, X is a known n × p matrix of rank p, β is a p × 1 unknown parameter vector and ε is an n × 1 vector of uncorrelated random variables with mean zero and variance σ<sup>2</sup> (i.e. we are not assuming that the ε<sub>i</sub> are normally distributed). Let β denote the least squares estimate of β, Ŷ denote the vector of fitted values, and let **R** be the vector of residuals. Find E(**R**) and cov(**R**). Show that cov(**R**, β) = 0 and cov(**R**, Ŷ) = 0.
- 10. For the simple linear regression model  $Y_i = a + bx_i + \varepsilon_i$ , i = 1, ..., n, where  $\sum_i x_i = 0$ and where the  $\varepsilon_i$  are iid  $N(0, \sigma^2)$  random variables, the MLEs  $\hat{a}$  and  $\hat{b}$  were found in Question 5. Find the distribution of  $\hat{\beta} = (\hat{a}, \hat{b})^T$ . Find a 95% confidence interval for b and for the mean value of Y when x = 1. [Hint: Look at "Applications of the distribution theory" in lectures.]
- +11 Consider the one-way ANOVA model of Question 8. Letting  $\bar{Y}_{i+} = n_i^{-1} \sum_{j=1}^{n_i} Y_{ij}$  and  $\bar{Y}_{i+} = n^{-1} \sum_{i=1}^{I} \sum_{j=1}^{n_i} Y_{ij}$  with  $n = n_1 + \ldots + n_I$ , show from first principles that the size  $\alpha$  likelihood ratio test of equality of means rejects  $H_0$  if

$$F \equiv \frac{\frac{1}{I-1} \sum_{i=1}^{I} n_i (\bar{Y}_{i+} - \bar{Y}_{i+})^2}{\frac{1}{n-I} \sum_{i=1}^{I} \sum_{j=1}^{n_i} (Y_{ij} - \bar{Y}_{i+})^2} > F_{I-1,n-I}(\alpha),$$

i.e. if 'the ratio of the between groups sum of squares to the within groups sum of squares is large'.