## Statistics: Example Sheet 3 (of 3)

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1. (a) Let $\mathbf{X} \sim N_{n}(\boldsymbol{\mu}, \Sigma)$, and let $A$ be an arbitrary $m \times n$ matrix. Prove directly from the definition that $A \mathbf{X}$ has an $m$-variate normal distribution. Show that $\operatorname{cov}(A \mathbf{X})=A \Sigma A^{T}$, and that $A \mathbf{X} \sim N_{m}\left(A \boldsymbol{\mu}, A \Sigma A^{T}\right)$. Give an alternative proof that $A \mathbf{X} \sim N_{m}\left(A \boldsymbol{\mu}, A \Sigma A^{T}\right)$ using moment generating functions.
(b) Let $\mathbf{X} \sim N_{n}(\boldsymbol{\mu}, \Sigma)$, and let $\mathbf{X}_{\mathbf{1}}$ denote the first $n_{1}$ components of $\mathbf{X}$. Let $\boldsymbol{\mu}_{\mathbf{1}}$ denote the first $n_{1}$ components of $\boldsymbol{\mu}$, and let $\Sigma_{11}$ denote the upper left $n_{1} \times n_{1}$ block of $\Sigma$. Show that $\mathbf{X}_{\mathbf{1}} \sim N_{n_{1}}\left(\boldsymbol{\mu}_{1}, \Sigma_{11}\right)$.
2. Let $X_{1}, \ldots, X_{n} \stackrel{i i d}{\sim} N\left(\mu, \sigma^{2}\right)$, where $\sigma^{2}$ is unknown, and suppose we are interested in testing $H_{0}: \mu=\mu_{0}$ against $H_{1}: \mu \neq \mu_{0}$. Letting $\bar{X}=n^{-1} \sum_{i=1}^{n} X_{i}$ and $S_{X X}=$ $\sum_{i=1}^{n}\left(X_{i}-\bar{X}\right)^{2}$, show that the likelihood ratio can be expressed as

$$
\Lambda_{\mathbf{X}}\left(H_{0}, H_{1}\right)=\left(1+\frac{T^{2}}{n-1}\right)^{n / 2}
$$

where $T=\frac{n^{1 / 2}\left(\bar{X}-\mu_{0}\right)}{\left\{S_{X X} /(n-1)\right\}^{1 / 2}}$. Determine the distribution of $T$ under $H_{0}$, and hence determine the size $\alpha$ likelihood ratio test.
3. Statisticians A and B obtain independent samples $X_{1}, \ldots, X_{10}$ and $Y_{1}, \ldots, Y_{17}$ respectively, both from a $N\left(\mu, \sigma^{2}\right)$ distribution with both $\mu$ and $\sigma^{2}$ unknown. They estimate $\left(\mu, \sigma^{2}\right)$ by $\left(\bar{X}, S_{X X} / 9\right)$ and ( $\bar{Y}, S_{Y Y} / 16$ ) respectively, where, for example, $\bar{X}=\frac{1}{10} \sum_{i=1}^{10} X_{i}$ and $S_{X X}=\sum_{i=1}^{10}\left(X_{i}-\bar{X}\right)^{2}$. Given that $\bar{X}=5.5$ and $\bar{Y}=5.8$, which statistician's estimate of $\sigma^{2}$ is more probable to have exceeded the true value by more than $50 \%$ ? Find this probability (approximately) in each case. [Hint: This is something of a 'trick' question. Why? You may find $\chi^{2}$ tables helpful.]
4. Suppose that $X_{1}, \ldots, X_{m}$ are iid $N\left(\mu_{X}, \sigma_{X}^{2}\right)$, and, independently, $Y_{1}, \ldots, Y_{n}$ are iid $N\left(\mu_{Y}, \sigma_{Y}^{2}\right)$, with $\mu_{X}, \mu_{Y}, \sigma_{X}^{2}$ and $\sigma_{Y}^{2}$ unknown. Write down the distributions of $S_{X X} / \sigma_{X}^{2}$ and $S_{Y Y} / \sigma_{Y}^{2}$. Derive a $100(1-\alpha) \%$ confidence interval for $\sigma_{X}^{2} / \sigma_{Y}^{2}$.
5. Consider the simple linear regression model $Y_{i}=a+b x_{i}+\varepsilon_{i}, i=1, \ldots, n$, where $\varepsilon_{1}, \ldots, \varepsilon_{n} \stackrel{i i d}{\sim} N\left(0, \sigma^{2}\right)$ and $\sum_{i=1}^{n} x_{i}=0$. Derive from first principles explicit expressions for the MLEs $\hat{a}, \hat{b}$ and $\hat{\sigma}^{2}$. Show that we can obtain the same expressions if we regard the simple linear regression model as a special case of the general linear model $\mathbf{Y}=X \boldsymbol{\beta}+\boldsymbol{\epsilon}$ and specialise the formulae $\hat{\boldsymbol{\beta}}=\left(X^{T} X\right)^{-1} X^{T} \mathbf{Y}$ and $\hat{\sigma}^{2}=n^{-1}\|\mathbf{Y}-X \hat{\boldsymbol{\beta}}\|^{2}$.
6. Consider the model $Y_{i}=b x_{i}+\varepsilon_{i}, i=1, \ldots, n$, where the $\varepsilon_{i}$ are independent with mean zero and variance $\sigma^{2}$ (regression through the origin). Write this in the form $\mathbf{Y}=X \boldsymbol{\beta}+\varepsilon$, and find the least squares estimator of $b$.

The relationship between the range in metres, $Y$, of a howitzer with muzzle velocity $v$ metres per second fired at angle of elevation $\alpha$ degrees is assumed to be $Y=$ $\frac{v^{2}}{g} \sin (2 \alpha)+\varepsilon$, where $g=9.81$ and where $\varepsilon$ has mean zero and variance $\sigma^{2}$. Estimate $v$ from the following independent observations made on 9 shells.

| $\alpha(\mathrm{deg})$ | 5 | 10 | 15 | 20 | 25 | 30 | 35 | 40 | 45 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\sin 2 \alpha$ | 0.1736 | 0.3420 | 0.5 | 0.6428 | 0.7660 | 0.8660 | 0.9397 | 0.9848 | 1 |
| range (m) | 4860 | 9580 | 14080 | 18100 | 21550 | 24350 | 26400 | 27700 | 28300 |

7. Consider the model $Y_{i}=\mu+\varepsilon_{i}, i=1, \ldots, n$, where $\varepsilon_{i}$ are iid $N\left(0, \sigma^{2}\right)$ random variables. Write this in matrix form $\mathbf{Y}=X \boldsymbol{\beta}+\boldsymbol{\varepsilon}$, and find the MLE $\hat{\boldsymbol{\beta}}$. Find the fitted values, the residuals and the residual sum of squares. Show how applying Theorem 13.2 (in lectures) to this case gives the independence of $\bar{Y}$ and $S_{Y Y}$ for an iid sample from $N\left(\mu, \sigma^{2}\right)$. Write down an unbiased estimate $\tilde{\sigma}^{2}$ of $\sigma^{2}$.
8. Consider the one-way analysis of variance (ANOVA) model $Y_{i j}=\mu_{i}+\varepsilon_{i j}, i=1, \ldots, I$, $j=1, \ldots, n_{i}$, where $\left(\varepsilon_{i j}\right) \stackrel{i i d}{\sim} N\left(0, \sigma^{2}\right)$. Derive from first principles explicit expressions for the MLEs $\hat{\mu}_{1}, \ldots, \hat{\mu}_{I}$ and $\hat{\sigma}^{2}$. Show that we can obtain the same expressions if we regard the ANOVA model as a special case of the general linear model $\mathbf{Y}=X \boldsymbol{\beta}+\boldsymbol{\varepsilon}$ and specialise the formulae $\hat{\boldsymbol{\beta}}=\left(X^{T} X\right)^{-1} X^{T} \mathbf{Y}$ and $\hat{\sigma}^{2}=n^{-1}\|\mathbf{Y}-X \hat{\boldsymbol{\beta}}\|^{2}$.
9. Consider the linear model $\mathbf{Y}=X \boldsymbol{\beta}+\boldsymbol{\varepsilon}$ where $\mathbf{Y}$ is an $n \times 1$ vector of observations, $X$ is a known $n \times p$ matrix of $\operatorname{rank} p, \boldsymbol{\beta}$ is a $p \times 1$ unknown parameter vector and $\boldsymbol{\varepsilon}$ is an $n \times 1$ vector of uncorrelated random variables with mean zero and variance $\sigma^{2}$ (i.e. we are not assuming that the $\varepsilon_{i}$ are normally distributed). Let $\hat{\boldsymbol{\beta}}$ denote the least squares estimate of $\boldsymbol{\beta}, \hat{\mathbf{Y}}$ denote the vector of fitted values, and let $\mathbf{R}$ be the vector of residuals. Find $\mathbb{E}(\mathbf{R})$ and $\operatorname{cov}(\mathbf{R})$. Show that $\operatorname{cov}(\mathbf{R}, \hat{\boldsymbol{\beta}})=0$ and $\operatorname{cov}(\mathbf{R}, \hat{\mathbf{Y}})=0$.
10. For the simple linear regresssion model $Y_{i}=a+b x_{i}+\varepsilon_{i}, i=1, \ldots, n$, where $\sum_{i} x_{i}=0$ and where the $\varepsilon_{i}$ are iid $N\left(0, \sigma^{2}\right)$ random variables, the MLEs $\hat{a}$ and $\hat{b}$ were found in Question 5. Find the distribution of $\hat{\boldsymbol{\beta}}=(\hat{a}, \hat{b})^{T}$. Find a $95 \%$ confidence interval for $b$ and for the mean value of $Y$ when $x=1$. [Hint: Look at "Applications of the distribution theory" in lectures.]
+11 Consider the one-way ANOVA model of Question 8. Letting $\bar{Y}_{i+}=n_{i}^{-1} \sum_{j=1}^{n_{i}} Y_{i j}$ and $\bar{Y}_{++}=n^{-1} \sum_{i=1}^{I} \sum_{j=1}^{n_{i}} Y_{i j}$ with $n=n_{1}+\ldots+n_{I}$, show from first principles that the size $\alpha$ likelihood ratio test of equality of means rejects $H_{0}$ if

$$
F \equiv \frac{\frac{1}{I-1} \sum_{i=1}^{I} n_{i}\left(\bar{Y}_{i+}-\bar{Y}_{++}\right)^{2}}{\frac{1}{n-I} \sum_{i=1}^{I} \sum_{j=1}^{n_{i}}\left(Y_{i j}-\bar{Y}_{i+}\right)^{2}}>F_{I-1, n-I}(\alpha),
$$

i.e. if 'the ratio of the between groups sum of squares to the within groups sum of squares is large'.

