

## Statistics: Example Sheet 3 (of 3)

Comments and corrections to [s.pitts@statslab.cam.ac.uk](mailto:s.pitts@statslab.cam.ac.uk)

1. If  $X \sim N(0, 1)$  and  $Y \sim \chi_n^2$  are independent, we say that  $T = \frac{X}{\sqrt{Y/n}}$  has a  $t$ -distribution with  $n$  degrees of freedom and write  $T \sim t_n$ . Show that the probability density function of  $T$  is  $f_T(t) = \frac{\Gamma(\frac{n+1}{2})}{\Gamma(\frac{n}{2})} \frac{1}{(n\pi)^{1/2}} \frac{1}{(1+t^2/n)^{(n+1)/2}}$ ,  $t \in \mathbb{R}$ .
2. Let  $X_1, \dots, X_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$ , where  $\sigma^2$  is unknown, and suppose we are interested in testing  $H_0 : \mu = \mu_0$  against  $H_1 : \mu \neq \mu_0$ . Letting  $\bar{X} = n^{-1} \sum_{i=1}^n X_i$  and  $S_{XX} = \sum_{i=1}^n (X_i - \bar{X})^2$ , show that the likelihood ratio can be expressed as

$$\Lambda_{\mathbf{X}}(H_0, H_1) = \left(1 + \frac{T^2}{n-1}\right)^{n/2},$$

where  $T = \frac{n^{1/2}(\bar{X} - \mu_0)}{\{S_{XX}/(n-1)\}^{1/2}}$ . Determine the distribution of  $T$  under  $H_0$ , and hence determine the size  $\alpha$  likelihood ratio test.

3. Suppose that  $X_1, \dots, X_m$  and  $Y_1, \dots, Y_n$  are independent, with  $X_i \sim N(\mu_X, \sigma^2)$ ,  $i = 1, \dots, m$ , and  $Y_i \sim N(\mu_Y, \sigma^2)$ ,  $i = 1, \dots, n$ , where  $\mu_X$ ,  $\mu_Y$  and  $\sigma^2$  are unknown. Write down the distributions of  $\bar{X} - \bar{Y}$  and  $\frac{S_{XX} + S_{YY}}{\sigma^2}$ . Find a  $100(1 - \alpha)\%$  confidence interval for  $\mu_X - \mu_Y$ .
4. Statisticians A and B obtain independent samples  $X_1, \dots, X_{10}$  and  $Y_1, \dots, Y_{17}$  respectively, both from a  $N(\mu, \sigma^2)$  distribution with both  $\mu$  and  $\sigma^2$  unknown. They estimate  $(\mu, \sigma^2)$  by  $(\bar{X}, S_{XX}/9)$  and  $(\bar{Y}, S_{YY}/16)$  respectively, where, for example,  $\bar{X} = \frac{1}{10} \sum_{i=1}^{10} X_i$  and  $S_{XX} = \sum_{i=1}^{10} (X_i - \bar{X})^2$ . Given that  $\bar{X} = 5.5$  and  $\bar{Y} = 5.8$ , which statistician's estimate of  $\sigma^2$  is more probable to have exceeded the true value by more than 50%? Find this probability (approximately) in each case. [Hint: This is something of a 'trick' question. Why? You may find  $\chi^2$  tables helpful.]
5. Suppose that  $X_1, \dots, X_m$  are iid  $N(\mu_X, \sigma_X^2)$ , and, independently,  $Y_1, \dots, Y_n$  are iid  $N(\mu_Y, \sigma_Y^2)$ , with  $\mu_X$ ,  $\mu_Y$ ,  $\sigma_X^2$  and  $\sigma_Y^2$  unknown. Write down the distributions of  $S_{XX}/\sigma_X^2$  and  $S_{YY}/\sigma_Y^2$ . Derive a  $100(1 - \alpha)\%$  confidence interval for  $\sigma_X^2/\sigma_Y^2$ .
6. Consider the simple linear regression model  $Y_i = a + bx_i + \varepsilon_i$ ,  $i = 1, \dots, n$ , where  $\varepsilon_1, \dots, \varepsilon_n \stackrel{iid}{\sim} N(0, \sigma^2)$  and  $\sum_{i=1}^n x_i = 0$ . Derive from first principles explicit expressions for the MLEs  $\hat{a}$ ,  $\hat{b}$  and  $\hat{\sigma}^2$ . Show that we can obtain the same expressions if we regard the simple linear regression model as a special case of the general linear model  $\mathbf{Y} = X\boldsymbol{\beta} + \boldsymbol{\varepsilon}$  and specialise the formulae  $\hat{\boldsymbol{\beta}} = (X^T X)^{-1} X^T \mathbf{Y}$  and  $\hat{\sigma}^2 = n^{-1} \|\mathbf{Y} - X\hat{\boldsymbol{\beta}}\|^2$ .

7. Consider the model  $Y_i = bx_i + \varepsilon_i$ ,  $i = 1, \dots, n$ , where the  $\varepsilon_i$  are independent with mean zero and variance  $\sigma^2$  (regression through the origin). Write this in the form  $\mathbf{Y} = X\boldsymbol{\beta} + \boldsymbol{\varepsilon}$ , and find the least squares estimator of  $b$ .

The relationship between the range in metres,  $Y$ , of a howitzer with muzzle velocity  $v$  metres per second fired at angle of elevation  $\alpha$  degrees is assumed to be  $Y = \frac{v^2}{g} \sin(2\alpha) + \varepsilon$ , where  $g = 9.81$  and where  $\varepsilon$  has mean zero and variance  $\sigma^2$ . Estimate  $v$  from the following independent observations made on 9 shells.

$\alpha$ (deg)	5	10	15	20	25	30	35	40	45
$\sin 2\alpha$	0.1736	0.3420	0.5	0.6428	0.7660	0.8660	0.9397	0.9848	1
range (m)	4860	9580	14080	18100	21550	24350	26400	27700	28300

8. Consider the model  $Y_i = \mu + \varepsilon_i$ ,  $i = 1, \dots, n$ , where  $\varepsilon_i$  are iid  $N(0, \sigma^2)$  random variables. Write this in matrix form  $\mathbf{Y} = X\boldsymbol{\beta} + \boldsymbol{\varepsilon}$ , and find the MLE  $\hat{\boldsymbol{\beta}}$ . Find the fitted values, the residuals and the residual sum of squares. Show how applying Theorem 3.7 (in lectures) to this case gives the independence of  $\bar{Y}$  and  $S_{YY}$  for an iid sample from  $N(\mu, \sigma^2)$ . Write down an unbiased estimate  $\tilde{\sigma}^2$  of  $\sigma^2$ .
9. Consider the one-way analysis of variance (ANOVA) model  $Y_{ij} = \mu_i + \varepsilon_{ij}$ ,  $i = 1, \dots, I$ ,  $j = 1, \dots, n_i$ , where  $(\varepsilon_{ij}) \stackrel{iid}{\sim} N(0, \sigma^2)$ . Derive from first principles explicit expressions for the MLEs  $\hat{\mu}_1, \dots, \hat{\mu}_I$  and  $\hat{\sigma}^2$ . Show that we can obtain the same expressions if we regard the ANOVA model as a special case of the general linear model  $\mathbf{Y} = X\boldsymbol{\beta} + \boldsymbol{\varepsilon}$  and specialise the formulae  $\hat{\boldsymbol{\beta}} = (X^T X)^{-1} X^T \mathbf{Y}$  and  $\hat{\sigma}^2 = n^{-1} \|\mathbf{Y} - X\hat{\boldsymbol{\beta}}\|^2$ .
10. Consider the linear model  $\mathbf{Y} = X\boldsymbol{\beta} + \boldsymbol{\varepsilon}$  where  $\mathbf{Y}$  is an  $n \times 1$  vector of observations,  $X$  is a known  $n \times p$  matrix of rank  $p$ ,  $\boldsymbol{\beta}$  is a  $p \times 1$  unknown parameter vector and  $\boldsymbol{\varepsilon}$  is an  $n \times 1$  vector of uncorrelated random variables with mean zero and variance  $\sigma^2$  (i.e. we are *not* assuming that the  $\varepsilon_i$  are normally distributed). Let  $\hat{\boldsymbol{\beta}}$  denote the least squares estimate of  $\boldsymbol{\beta}$ ,  $\hat{\mathbf{Y}}$  denote the vector of fitted values, and let  $\mathbf{R}$  be the vector of residuals. Find  $\mathbb{E}(\mathbf{R})$  and  $\text{cov}(\mathbf{R})$ . Show that  $\text{cov}(\mathbf{R}, \hat{\boldsymbol{\beta}}) = 0$  and  $\text{cov}(\mathbf{R}, \hat{\mathbf{Y}}) = 0$ .
11. For the simple linear regression model  $Y_i = a + bx_i + \varepsilon_i$ ,  $i = 1, \dots, n$ , where  $\sum_i x_i = 0$  and where the  $\varepsilon_i$  are iid  $N(0, \sigma^2)$  random variables, the MLEs  $\hat{a}$  and  $\hat{b}$  were found in Question 6. Find the distribution of  $\hat{\boldsymbol{\beta}} = (\hat{a}, \hat{b})^T$ . Find a 95% confidence interval for  $b$  and for the mean value of  $Y$  when  $x = 1$ . [Hint: Look at ‘‘Applications of the distribution theory’’ in lectures.]
- +12 Consider the one-way ANOVA model of Question 9. Letting  $\bar{Y}_{i+} = n_i^{-1} \sum_{j=1}^{n_i} Y_{ij}$  and  $\bar{Y}_{++} = n^{-1} \sum_{i=1}^I \sum_{j=1}^{n_i} Y_{ij}$  with  $n = n_1 + \dots + n_I$ , show from first principles that the size  $\alpha$  likelihood ratio test of equality of means rejects  $H_0$  if

$$F \equiv \frac{\frac{1}{I-1} \sum_{i=1}^I n_i (\bar{Y}_{i+} - \bar{Y}_{++})^2}{\frac{1}{n-I} \sum_{i=1}^I \sum_{j=1}^{n_i} (Y_{ij} - \bar{Y}_{i+})^2} > F_{I-1, n-I}(\alpha),$$

i.e. if ‘the ratio of the between groups sum of squares to the within groups sum of squares is large’.