

## Some common discrete distributions

Distribution	Notation	pmf $f(x)$	Range of $X$	Parameter range	$\mathbb{E}(X)$	$\text{Var}(X)$	$\mathbb{E}(z^X)$
Discrete uniform	$X \sim U\{1, \dots, n\}$	$\frac{1}{n}$	$\{1, \dots, n\}$	$n \in \mathbb{N}$	$\frac{1}{2}(n+1)$	$\frac{1}{12}(n^2-1)$	$\frac{1}{n} \sum_{i=1}^n z^i$
Binomial	$X \sim \text{Bin}(n, p)$	$\binom{n}{x} p^x (1-p)^{n-x}$	$\{0, 1, \dots, n\}$	$n \in \mathbb{N}, p \in [0, 1]$	$np$	$np(1-p)$	$\{pz + (1-p)\}^n$
Poisson	$X \sim \text{Poisson}(\lambda)$	$e^{-\lambda} \frac{\lambda^x}{x!}$	$\{0, 1, \dots\}$	$\lambda \in [0, \infty)$	$\lambda$	$\lambda$	$e^{\lambda(z-1)}$
Negative binomial	$X \sim \text{NegBin}(k, p)$	$\binom{x-1}{k-1} p^k (1-p)^{x-k}$	$\{k, k+1, \dots\}$	$k \in \mathbb{N}, p \in [0, 1]$	$\frac{k}{p}$	$\frac{k(1-p)}{p^2}$	$\frac{(pz)^k}{\{1-(1-p)z\}^k}$
Multinomial	$X \sim \text{Multi}(n, p_1, \dots, p_k)$	$\frac{n!}{x_1! \dots x_k!} p_1^{x_1} \dots p_k^{x_k}$	$(n_1, \dots, n_k) \in \{0, 1, \dots, n\}^k : \sum_i n_i = n$	$p_1, \dots, p_k \in [0, 1] : \sum_i p_i = 1, n \in \mathbb{N}$	$(np_1, \dots, np_k)$	$\text{Cov}(X_i, X_j) = \begin{cases} np_i(1-p_i) & i=j \\ -np_i p_j & i \neq j \end{cases}$	$\mathbb{E}(z_1^{X_1} \dots z_k^{X_k}) = (\sum_{i=1}^k p_i z_i)^n$

### Notes:

1. The  $\text{Bin}(1, p)$  distribution is also called the Bernoulli( $p$ ) distribution. If  $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Bernoulli}(p)$ , then  $\sum_{i=1}^n X_i \sim \text{Bin}(n, p)$ . The  $\text{Bin}(n, p)$  distribution models the number of successes in  $n$  independent trials, each with probability  $p$  of success.
2. The  $\text{NegBin}(1, p)$  distribution is also called the Geometric( $p$ ) distribution. If  $X_1, \dots, X_k \stackrel{iid}{\sim} \text{Geometric}(p)$ , then  $\sum_{i=1}^k X_i \sim \text{NegBin}(k, p)$ . The  $\text{NegBin}(k, p)$  distribution models the number of independent trials required to attain  $k$  successes, each with probability  $p$  of success.
3. The  $\text{Multi}(n, p_1, \dots, p_k)$  distribution models the number of balls that appear in each of  $k$  buckets, when  $n$  balls are placed independently in the buckets and a ball falls in the  $i$ th bucket with probability  $p_i$ .

## Some common (absolutely) continuous distributions

Distribution	Notation	pdf $f(x)$	Range	Parameter range	$\mathbb{E}(X)$	$\text{Var}(X)$	$\mathbb{E}(e^{tX})$
Uniform	$X \sim U[a, b]$	$\frac{1}{b-a}$	$[a, b]$	$(a, b) \in \mathbb{R}^2, a < b$	$\frac{1}{2}(a+b)$	$\frac{1}{12}(b-a)^2$	$\frac{e^{bt}-e^{at}}{t(b-a)}$
Normal	$X \sim N(\mu, \sigma^2)$	$\frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(x-\mu)^2}$	$\mathbb{R}$	$\mu \in \mathbb{R}, \sigma \in (0, \infty)$	$\mu$	$\sigma^2$	$e^{t\mu + \sigma^2 t^2 / 2}$
Gamma	$X \sim \text{Gamma}(\alpha, \lambda)$	$\frac{\lambda^\alpha x^{\alpha-1} e^{-\lambda x}}{\Gamma(\alpha)}$	$(0, \infty)$	$\alpha \in (0, \infty), \lambda \in (0, \infty)$	$\frac{\alpha}{\lambda}$	$\frac{\alpha}{\lambda^2}$	$\begin{cases} (\frac{\lambda}{\lambda-t})^\alpha & \text{if } t < \lambda \\ \infty & \text{if } t \geq \lambda \end{cases}$
Beta	$X \sim \text{Beta}(a, b)$	$\frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} x^{a-1} (1-x)^{b-1}$	$(0, 1)$	$a \in (0, \infty), b \in (0, \infty)$	$\frac{a}{a+b}$	$\frac{ab}{(a+b)^2(a+b+1)}$	
Cauchy	$X \sim \text{Cauchy}$	$\frac{1}{\pi(1+x^2)}$	$\mathbb{R}$	Does not exist	$\infty$	$\infty$	$\begin{cases} 1 & \text{if } t = 0 \\ \infty & \text{if } t \neq 0 \end{cases}$
Multivariate normal	$X \sim N_d(\mu, \Sigma)$	$\frac{e^{-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)}}{(2\pi)^{d/2} (\det \Sigma)^{1/2}}$	$\mathbb{R}^d$	$\mu \in \mathbb{R}^d, \Sigma$ pos. def.	$\mu$	$\text{Cov}(X_i, X_j) = \Sigma_{ij}$	$\mathbb{E}(e^{t^T X}) = e^{t^T \mu + t^T \Sigma t / 2}$

### Notes:

1. The Gamma(1,  $\lambda$ ) distribution is the same as the Exponential( $\lambda$ ) distribution. If  $X_1, \dots, X_n \stackrel{iid}{\sim}$  Exponential( $\lambda$ ), then  $\sum_{i=1}^n X_i \sim \text{Gamma}(n, \lambda)$ .
2. For  $n \in \mathbb{N}$ , the Gamma( $\frac{n}{2}, \frac{1}{2}$ ) distribution is the same as the  $\chi_n^2$  distribution. If  $X_1, \dots, X_n \stackrel{iid}{\sim} N(0, 1)$ , then  $\sum_{i=1}^n X_i^2 \sim \chi_n^2$ . If  $Y \sim \text{Gamma}(n, \lambda)$  then  $2\lambda Y \sim \chi_{2n}^2$ .

3. Recall that the Gamma function is defined, for  $z \in \mathbb{C}$  with  $\text{Re}(z) > 0$ , by  $\Gamma(z) = \int_0^\infty t^{z-1} e^{-t} dt$ . If  $n \in \mathbb{N}$ , then  $\Gamma(n) = (n-1)!$ . The function  $B(a, b) = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}$  is called the beta function.

4. More generally, we can define the degenerate normal distribution: say  $X \sim N(\mu, 0)$  if  $\mathbb{P}(X = \mu) = 1$ . Then we say  $X = (X_1, \dots, X_d) \sim N_d(\mu, \Sigma)$  if every linear combination  $t_1 X_1 + \dots + t_d X_d$  has a (possibly degenerate) univariate normal distribution. This more general definition includes situations like the following: suppose  $X_1 \sim N(0, 1)$ , and let  $X = (X_1, X_1)$ . Then  $X \sim N_2(0, \Sigma)$ , where  $\Sigma = \begin{pmatrix} 1 & 1 \\ 1 & 1 \end{pmatrix}$ . Note here that  $\det \Sigma = 0$ .